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Anion gap corrected for albumin, phosphate and lactate is a good predictor of strong ion gap in critically ill patients: a nested cohort study

Ânion gap corrigido para albumina, fosfato e lactato é um bom preditor de íon gap forte em pacientes enfermos graves: estudo de coorte em nicho

ABSTRACT

Objective: Corrected anion gap and strong ion gap are commonly used to estimate unmeasured anions. We evaluated the performance of the anion gap corrected for albumin, phosphate and lactate in predicting strong ion gap in a mixed population of critically ill patients. We hypothesized that anion gap corrected for albumin, phosphate and lactate would be a good predictor of strong ion gap, independent of the presence of metabolic acidosis. In addition, we evaluated the impact of strong ion gap at admission on hospital mortality.

Methods: We included 84 critically ill patients. Correlation and agreement between the anion gap corrected for albumin, phosphate and lactate and strong ion gap was evaluated by the Pearson correlation test, linear regression, a Bland-Altman plot and calculating interclass correlation coefficient. Two subgroup analyses were performed: one in patients with base-excess <-2mEq/L (low BE group - lBE) and the other in patients with base-excess >-2mEq/L (high BE group - hBE). A logistic

regression was performed to evaluate the association between admission strong ion gap levels and hospital mortality.

Results: There was a very strong correlation and a good agreement between anion gap corrected for albumin, phosphate and lactate and strong ion gap in the general population (r²=0.94; bias 1.40; limits of agreement -0.75 to 3.57). Correlation was also high in the IBE group (r²=0.94) and in the hBE group (r²=0.92). High levels of strong ion gap were present in 66% of the whole population and 42% of the cases in the hBE group. Strong ion gap was not associated with hospital mortality by logistic regression.

Conclusion: Anion gap corrected for albumin, phosphate and lactate and strong ion gap have an excellent correlation. Unmeasured anions are frequently elevated in critically ill patients with normal base-excess. However, there was no association between unmeasured anions and hospital mortality.

Keywords: Acid-base equilibrium/ physiology; Critical illness; Blood chemical analysis; Serum albumin/blood; Phosphate/ blood; Lactate/blood; Prognosis

INTRODUCTION

Evaluating the acid-base status of a critically ill patient is crucial, as it implies both prognostic and therapeutic considerations. (1-5) Elevated unmeasured anions are a common cause of metabolic acidosis in such patients and may be related to prognosis. (3,6) The two most common ways to estimate unmeasured anions are

the anion gap (AG) calculation and the physicochemical approach of the strong ion gap (SIG) calculation, as described by Stewart. (4.7)

There are some caveats in the interpretation of the AG. AG is usually calculated as ([Na] + [K] - [HCO₃ + [Cl]), ignoring the effects of other relevant compounds, such as albumin, lactate and phosphate. (7) To refine the meaning of AG as a surrogate for unmeasured anions, AG can be adjusted for such ions. (3) The calculation of SIG, on the other hand, is cumbersome and not practical at the bedside, although it yields results that are believed to be better estimates of unmeasured anions. (8)

The use of AG to predict SIG has been performed in other studies, (9-12) mostly in acidotic patients, (9,11) but no studies have reported the prediction of SIG using phosphate-corrected AG. Moreover, there are also few reports on the presence of unmeasured anions in the absence of "measured" metabolic acidosis (i.e., normal base excess - BE),(5,11) and the exact impact of unmeasured anions on prognosis is unclear. (13-15) Therefore, this study aimed to evaluate the performance of the anion gap corrected for albumin, phosphate and lactate (AGCAPL) to predict SIG in a mixed population of critically ill patients. We hypothesized that AGCAPL would be a good predictor of SIG and that the prediction capability would be independent of the presence of metabolic acidosis, as evaluated through base excess. As a secondary objective, we also evaluated the association between admission SIG values and hospital mortality using logistic regression.

METHODS

This study is a nested prospective cohort study that included 84 critically ill patients who were admitted to the intensive care unit (ICU) from January 2011 to December 2012. We included patients who took part in a prospective observational study that aimed to evaluate inflammatory markers in critical illness. (16) After informed consent was obtained from the patient or legal representative, blood was collected during ICU admission. General laboratory data were processed, and the results were stored on a database. Patients were followed until death or hospital discharge using hospital's online system. The study was approved by the local Hospital das Clínicas ethics committee (registration number 1.207/09).

AG was calculated as $[([Na] + [K] - [Cl] - [HCO_3])$. AGCAPL was calculated as AGCAPL = $[([Na] + [K] - [Cl] - [HCO_3])$ - $(2 \times albumin g/dL + 0.5 \times phosphate mg/dL)$ - [lactate mmol/L]], as previously shown. SIG was defined as the difference between apparent and effective strong ion difference (SIDa and SIDe, respectively). SIDa was calculated as SIDa = [Na] + [K] + [Mg] + [Ca] - [Cl] - [Lactate]. SIDe was calculated as SIDe = $12.2 \times pCO_2/(10^{-pH}) + 10 \times [albumin] \times (0.123 \times pH - 0.631) + [PO_4^-] \times (0.309 \times pH - 0.469)$. Blood gas analysis and lactate measurement were performed using an OMNI analyzer (Roche Diagnostics System, F. Hoffmann, La Roche Ltd, Basel, Switzerland).

Data were tested for normality using the Kolmogorov-Smirnoff or Shapiro-Wilkes test, as appropriate. Continuous normal data were compared using the t-test or analysis of variance. Continuous data that were not normal were compared using the Mann-Whitney test or Kruskal-Wallis test, as appropriate. Fisher's exact test or Chi-squared tests were used for dichotomous variables.

After confirming the normal distribution of the involved variables, we analyzed the correlation between AGCAPL and SIG using a Pearson correlation test. Agreement was evaluated through a Bland-Altman plot and the measurement of the interclass correlation coefficient (ICC). We also built a linear regression between AGCAPL and SIG and obtained the R^2 of the prediction model. Two subgroup analyses were performed: one in patients with low base excess (BE <-2mEq/L - lBE group) and the other in patients with high base excess (BE >-2mEq/L - hBE group).

To evaluate the association between admission SIG levels and hospital mortality, we built one logistic regression using hospital mortality as the outcome. Variables associated with a p<0.25 by univariate analysis were included in the analysis. A stepwise regression was performed to identify variables that were independently associated with hospital mortality. The prediction capability of SIG for hospital mortality was also evaluated through the creation of an ROC curve and the calculation of the area under the curve. (19)

All analyses were performed using the R software (www.r-project.org) with the pROC and car packages. Mountain plots were created using MedCalc version 12.7.0 (MedCalc Software, Acacialaan 22, B-8400 Ostend, Belgium). A p<0.05 was considered significant for all analyses.

RESULTS

The general characteristics of the patients who were included in the study are shown in table 1. Demographic data were similar between patients with metabolic acidosis (lBE) and patients without metabolic acidosis (hBE) at ICU admission. lBE patients had higher illness severity, as evaluated by the total SOFA score (7 [3-9] versus 3.5 [2-6.5]; p<0.01), but similar SAPS3 levels (55 [43.2-68.2] versus 51 [41.7-61.2]; p=0.51). Admission due to sepsis was more common in the lBE than in the hBE group (66% versus 30%; p<0.01). ICU and hospital mortality was similar for both groups. Patients in the lBE group had lower values for pH, BE, bicarbonate, CO₂, SIDa and albumin. The lBE group had higher levels of lactate, AG, AGCAPL and SIG.

There was a very strong correlation between AGCAPL and SIG in the general population (r^2 =0.94; Figure 1), and the ICC was high (0.93; CI 0.89-0.95). The bias was 1.40, and the limits of agreement were -0.75 to

3.57, as shown in the Bland-Altman plot (Figure 2). The correlation was also high in the IBE group (r^2 =0.94; ICC = 0.91, CI 0.85-0.95; bias = 1.55, limits of agreement from -0.67 to 3.75) and in the hBE group (r^2 =0.92; ICC = 0.90, CI 0.82-0.95; bias = 1.22; limits of agreement from -0.85 to 3.3).

Uncorrected AG had a much weaker correlation with SIG in the general population (r²=0.66), IBE group (r²=0.58) and hBE group (r²=0.67). A mountain plot for the difference between uncorrected AG, AGCALP and SIG for each percentile of SIG is shown in figure 3. Values of SIDa, SIDe and SIG for the whole population, IBE group, hBE group and hBE group with high SIG are shown in a violin plot in figure 4.

High levels of SIG (above 6mEq/L)⁽⁷⁾ were present in 56 patients (66% of the whole population). In the hBE group, 16 patients (42 %) had a SIG above 6mEq/L. The biochemical comparison between patients from the hBE group with high (>6mEq/L) and low SIG values is shown in table 2.

Table 1 - Clinical and laboratory features of studied patients

Features	All patients (N=84)	Low base-excess (N=48)	High base-excess (N=36)	p value
Age (years)	50.21 (17.40)	51.02 (18.70)	49.13 (15.64)	0.61
Sex, male	49 (58)	25 (52)	20 (55)	0.75
SAPS3	52.50 (41.75-64.75)	55 (43.25-68.25)	51 (41.75-61.25)	0.51
SOFA at admission	5 (2-9)	7 (3-9)	3.5 (2-6.5)	< 0.01
Sepsis	43 (51)	32 (66)	11 (30)	< 0.01
ICU mortality	20 (23)	13 (27)	7 (19)	0.41
Hospital mortality	26 (30)	18 (37)	8 (22)	0.13
рН	7.38 (7.33-7.41)	7.36 (7.29-7.40)	7.40 (7.38-7.41)	< 0.01
BE (mEq/L)	-2.4 (-4.87-0.75)	-4.55 (-8.47-3.2)	-0.3 (-1.12-2.65)	< 0.01
HCO ₃ - (mEq/L)	21.8 (19-24)	19.65 (17.45-21.50)	24.4 (23.28-27.22)	< 0.01
PCO ₂ (mmHg)	37.1 (32.45-44)	34.15 (29.98-39.78)	41 (37.32-46.02)	< 0.01
Na (mEq/L)	139 (136-143.2)	138 (135.8-143)	140 (137-144)	0.11
CI (mEq/L)	104 (104-108.2)	104.5 (101.8-108)	104 (100-109)	0.88
Mg (mg/dL)	1.95 (1.68-2.22)	1.97 (1.68-2.13)	1.93 (1.68-2.32)	0.73
Ca (mg/dL)	5 (4-5)	4.5 (4-5)	5.0 (4.75-5.0)	0.01
P (mg/dL)	3.4 (2.77-4.95)	3.65 (2.67-5.10)	3.30 (2.8-4.3)	0.57
Albumin (mg/dL)	3.0 (2.3-3.5)	2.8 (2.1-3.2)	3.35 (2.87-3.62)	< 0.01
Lactate (mmol/L)	1.56 (1.11-2.48)	1.83 (1.27-2.61)	1.38 (1.11-1.80)	0.03
AG (mEq/L)	16.50 (13.92-19.12)	17.35 (16.05-20.98)	15.15 (12.85-16.70)	< 0.01
AGCAPL (mEq/L)	6.69 (4.34-9.34)	8.68 (6.34-10.52)	4.95 (2.69-6.61)	< 0.01
SIG (mEq/L)	8.12 (5.79-10.92)	9.82 (7.76-12.32)	5.82 (3.73-7.41)	< 0.01
SIDa (mEq/L)	40.90 (37.13-43.86)	39.24 (36.37-41.31)	43.26 (40.26-45.74)	0.01
SIDe (mEq/L)	32.09 (28.45-35.90)	29.11 (26.61-31.18)	36.29 (34.10-39.40)	< 0.01

SAPS3 - Simplified Acute Physiology Score; SOFA - Sepsis Related Organ Failure Assessment; ICU - intensive care unit; BE - base-excess; AG - anion gap; AGCAPL - anion gap corrected for albumin, phosphate and lactate; SIG - strong ion gap; SIDa - apparent strong ion difference; SIDe - effective strong ion difference. p value for high versus low BE groups. The results are expressed as a number (%), mean or median (IQ) ± standard deviation.

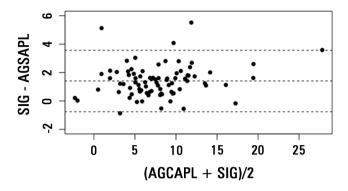


Figure 1 - Plot of anion gap corrected for albumin, phosphate and lactate (AGCAPL) versus strong ion gap (SIG) for the general population.

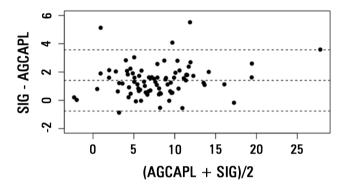


Figure 2 - Bland-Altman plot of anion gap corrected for albumin, phosphate and lactate (AGCAPL) and strong ion gap (SIG). The central dashed line represents bias (mean of SIG - AGCAPL). The other two dashed lines represent 95% limits of agreement, equivalent to 1.96 x SD of the differences (SIG - AGCAPL), equivalent to -0.75 and 3.57.

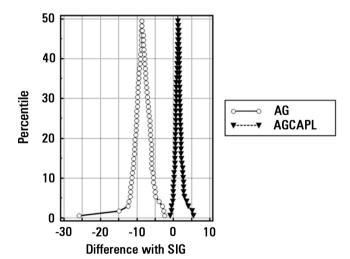


Figure 3 - Mountain plot for the comparison between anion gap; anion gap (AG) corrected for albumin, phosphate and lactate (AGCAPL); and strong ion gap (SIG) values. Note that there is a narrow correlation between AGCAPL and SIG, but not between AG and SIG.

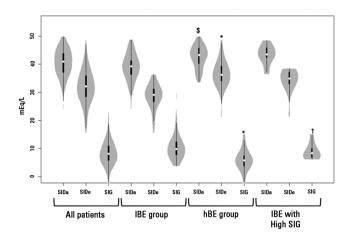


Figure 4 - Violin plot of SIDa, SIDe and strong ion gap for the general population, IBE group, hBE group and hBE patients with strong ion gap above 6mEq/L. White dots: median values; Black boxes: interquartiles; Black lines: ranges; gray shading: density of observed values. * $^{\circ}$ P=0.01 versus IBE group; * $^{\circ}$ P<0.01 versus IBE group; † $^{\circ}$ P<0.01 versus IBE group; * $^{\circ}$ P<0.03 versus IBE group; * $^{\circ}$ P<0.03 versus IBE group; * $^{\circ}$ P<0.04 versus IBE group; * $^{\circ}$ P<0.05 versus IBE group; * $^{\circ}$ P<0.05 versus IBE group; * $^{\circ}$ P<0.06 versus IBE group; * $^{\circ}$ P<0.07 versus IBE group; * $^{\circ}$ P<0.08 versus IBE group; * $^{\circ}$ P<0.09 versus IBE group; * $^{\circ}$ P<0.00 versus IBE group; * $^{\circ}$

Table 2 - Laboratory comparison between patients with high (>6mEq/L) and normal strong ion gap (<6mEq/L) values in the hBE group

	Low SIG (N=20)	High SIG (N=16)	p value
рН	7.39 (7.37-7.39)	7.40 (7.39-7.41)	0.70
PCO ₂ (mmHg)	43.4 (40.00-46.70)	37.60 (35.88-42.50)	0.01
Na (mEq/L)	141 (137-144.8)	139 (137.5-142.2)	0.42
CI (mEq/L)	106 (101.5-110.5)	103 (100-106.5)	0.17
Albumin (mg/dL)	3.5 (3.02-3.41)	3.1 (2.85-3.50)	0.09
Lactate (mmol/L)	1.38 (1.19-1.91)	1.33 (0.97-1.67)	0.35
SIG (mEq/L)	4.13 (2.75-5.20)	8.37 (6.85-10.02)	< 0.01
SIDa (mEq/L)	42.68 (38.83-45.74)	43.45 (42.38-45.79)	0.35
SIDe (mEq/L)	38.64 (34.75-40.97)	34.90 (33.00-37.21)	0.01

SIG - strong ion gap; SIDa - apparent strong ion difference; SIDe - effective strong ion difference. The results are expressed as the median $(IQ) \pm \text{standard deviation}$.

The variables included in the mortality prediction model were total SOFA score, SAPS3 score, albumin levels, pH, BE, lactate levels, SIG and diagnosis of sepsis. After stepwise regression, only SAPS3 (OR 1.04; CI 95% 1.01-1.08 per point increase) and albumin levels (OR 0.19; CI 95% 0.07-0.49 per point increase) were associated with hospital mortality. SIG had a poor prediction capability for hospital mortality (AUC 0.61; CI 95% 0.47-0.74).

DISCUSSION

Our analysis demonstrates that AGCAPL is highly correlated with SIG in a mixed sample of critically ill patients and that such correlation is independent of the presence of metabolic acidosis. AGCAPL performed much better than uncorrected AG in the general

population and subgroups, as seen by the Pearson correlation and mountain plot results. We have also shown that unmeasured anions are frequently elevated in critically ill patients, even in patients with an apparently normal BE (42% of cases). Additionally, we have shown that SIG is not associated with hospital mortality by logistic regression.

The agreement between corrected anion gap and SIG has been shown in other studies. Moviat et al. have previously evaluated the correlation of albumin- and lactate-corrected anion gap with SIG in acidotic (BE <-5mEq/L) patients. (9) They concluded that SIG and AG were strongly correlated (r²=0.93), with a small bias, (9) but the interclass coefficient correlation was not reported. We obtained values similar to those from Moviat et al. (9) for the correlation between AG and AGCAPL, suggesting that the addition of the correction for phosphate levels does not improve the correlation between AG and SIG. However, taking agreement into consideration, the bias between SIG and the corrected anion gap was slightly lower in our analysis (1.86 versus 1.40), which may have occurred because phosphate levels were taken into account or due to a difference in samples. Therefore, the only benefit of adding the correction for phosphate in the AG is a small reduction in the bias, which is most likely clinically irrelevant. Martin et al. (1) and Dubin et al. (12) also reported an excellent correlation between albumin-corrected AG and SIG. Finally, Abdulraof Menesi et al. have shown that both the traditional and the physicochemical approaches are similar, even in specific populations, such as patients with kidney graft. (10) Abdulraof Menesi et al. also suggested that the correlation between AG and SIG could be reduced when AG levels were low. (10) In our analysis, we also found a lower slope for the relationship between AGCAPL and SIG in patients with low AGCAPL (<10) compared to those with high AGCAPL (>10) (r=0.97 and r=1.07, respectively - data not shown). This finding was not the primary endpoint of our study and deservers further evaluation in larger samples.

The correct diagnosis of the metabolic disturbance is important for clinical management. High lactate levels are related to hypoperfusion, while hyperchloremia may be the result of aggressive resuscitation using chloride-rich solutions. (20-23) The latter consideration may be particularly important because hyperchloremia has been suggested to be associated with mortality. (24) The association between unmeasured anions and hypoperfusion, however, is less clear. It remains to be determined if SIG levels can be used as a marker of hypoperfusion or if it is appropriate to institute a specific treatment (e.g., fluid loading,

inotropes) when unmeasured anions are increased. Nevertheless, unmeasured anions are frequently elevated in critical illness, reportedly elevated in more than 90% of trauma patients when a low cutoff is used, (1) and may be associated with poor prognosis. (1,3,6) Patients may present with "occult" metabolic acidosis, i.e., the higher levels of unmeasured anions may be masked by concomitantly reduced albumin levels. (5) In our analysis, 42% (16) of the patients with normal BE had high levels of unmeasured anions. When those patients were compared with the remaining patients in the hBE group with low SIG (20 patients) (Table 2), they had a lower PCO, and lower SIDe. Therefore, the high SIG levels were compensated for by a reduction in SIDe that was mainly caused by decreased albumin (Table 2) and a reduction in PCO₂, highlighting the complex acid-base behavior in critical illness. (25) AGCAPL still had an excellent correlation with SIG in these patients ($r^2=0.95$).

The impact of SIG on mortality is also debated, with some reports showing an association with mortality, while others found no association. (13-15,26) Durward et al. showed that SIG was better than lactate as a predictor of mortality after cardiac surgery in children, although no model was built. (13) In trauma patients, Kaplan et al. reported that SIG was a good predictor of mortality and was the strongest factor associated with mortality by logistic regression. (26) On the other hand, Rocktaeschel et al. (15) found that, despite being associated with mortality by multivariate analysis, SIG values had a poor prediction capability for mortality. Recently, Ratanarat et al. also suggested that SIG was higher in non-survivors, although no multivariate model was implemented to reduce confounding variables. (27) In our analysis, SIG values were not associated with in-hospital mortality after logistic regression. In fact, not even BE was associated with mortality in our analysis, contrary to previous research on the subject. (28) The prediction capability of SIG for hospital mortality was also poor (AUC 0.61; 95% CI 0.47-0.74). This finding may be explained by the mix of clinical diagnosis at admission, particular sample features or both.

The best method for the interpretation of acid-base disorders at the bedside remains to be defined. (4) Dubin et al. have shown that the physicochemical approach offers no advantage over the traditional approach with corrected AG; in fact, the physicochemical approach allowed for an additional diagnosis of metabolic acidosis in only 1% of the cases. (12) Martin et al. have shown that the results of the traditional approach (using anion gap) and the physicochemical approach yield different clinical interpretations in up to 28% of trauma patients, (1) while

Kaplan et al. have suggested that a physicochemical approach can improve the accuracy of acid-base disturbances in trauma and, therefore, reduce inappropriate fluid loading due to the suspicion of hypoperfusion-induced metabolic acidosis. (29) Boniatti et al. reported a series of 175 patients and evaluated the percentage of cases in which the physicochemical approach would supply different results when compared to the traditional approach. (30) The authors concluded that the physicochemical approach would allow for an additional diagnosis of metabolic disorder (specifically, a decrease in SIDe) in over 33% of the cases. However, several caveats in the methods, such as the use of different thresholds for BE and SID and misinterpretations of the acid-base status, limit the validity of the data. (30,31) As Dubin et al. have previously stated, both methods will most likely yield similar results if they are properly applied. (31) Although this question has yet to be settled, our findings suggest that AGCAPL is a good surrogate for SIG measurements, with the advantage of being more easily calculated at the bedside.

Our study has several limitations. First, our small sample limits subgroup analysis and reduces external validity. Because this is a single-center study, the bias between SIG and AGCAPL reflects the evaluated population. Despite including a significant range of diagnoses, these results cannot be generalized to other specific populations in different

settings. Second, we did not evaluate the impact of any approach on clinical management. Therefore, the impact of the measurement of unmeasured anions at the bedside should be explored in further studies. Third, we only evaluated ICU admission values. Consequently, the impact of acid-base status changes during ICU stay on prognosis was not possible.

CONCLUSION

Unmeasured anions are frequently elevated in critically ill patients. Because there is a strong correlation between anion gap corrected for albumin, phosphate and lactate and strong ion gap in patients with both normal and low base-excess, anion gap corrected for albumin, phosphate and lactate may be used as a surrogate for strong ion gap at the bedside. Strong ion gap values at admission are not associated with in-hospital mortality.

Authors' contributions

FG Zampieri, LM Cruz Neto, HP Souza, FP Silva designed the study and conducted data collection. FG Zampieri, OT Ranzani, M Park and AT Maciel performed statistical analysis. FG Zampieri and M Park wrote the manuscript. All authors reviewed the manuscript in its final form.

RESUMO

Objetivo: Ânion *gap* corrigido e íon *gap* forte são usados comumente para estimar os ânions não medidos. Avaliamos o desempenho do ânion *gap* corrigido para albumina, fosfato e lactato na predição do íon *gap* forte em uma população mista de pacientes enfermos graves. Formulamos a hipótese de que o ânion *gap* corrigido para albumina, fosfato e lactato seria um bom preditor do íon *gap* forte, independentemente da presença de acidose metabólica. Além disso, avaliamos o impacto do íon *gap* forte por ocasião da admissão na mortalidade hospitalar.

Métodos: Incluímos 84 pacientes gravemente enfermos. A correlação e a concordância entre o ânion *gap* corrigido para albumina, fosfato e lactato e o íon *gap* forte foi avaliada utilizando-se os testes de correlação de Pearson, regressão linear, plot de Bland-Altman e pelo cálculo do coeficiente de correlação interclasse. Foram realizadas duas análises de subgrupos: uma para pacientes com excesso de base <-2mEq/L (grupo com baixo excesso de base) e outro grupo de pacientes com excesso de base >-2mEq/L (grupo com alto excesso de base). Foi realizada uma regressão logística para avaliar a associação entre os níveis de íon *gap* forte na admissão e a mortalidade hospitalar.

Resultados: Houve correlação muito forte e uma boa concordância entre o ânion *gap* corrigido para albumina, fosfato e lactato e o íon *gap* forte na população geral (r²=0,94; bias 1,40; limites de concordância de -0,75 a 3,57). A correlação foi também elevada nos grupos com baixo excesso de base (r²=0,94) e alto excesso de base (r²=0,92). Estavam presentes níveis elevados de íon *gap* forte em 66% da população total e 42% dos casos do grupo alto excesso de. Íon *gap* forte não se associou com a mortalidade hospitalar, conforme avaliação pela regressão logística.

Conclusão: O ânion *gap* corrigido para albumina, fosfato e lactato e o íon *gap* forte tiveram uma excelente correlação. Os ânions não medidos estão frequentemente elevados em pacientes gravemente enfermos com excesso de base normal. Entretanto, não ocorreu associação entre os ânions não medidos e a mortalidade hospitalar.

Descritores: Equilíbrio acido-base/fisiologia; Estado terminal; Análise química do sangue; Albumina sérica/sangue; Fosfatos/sangue; Lactatos/sangue; Prognóstico

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