

MIXED-EFFECTS GROWTH CURVES IN THE EVALUATION OF NELLORE SIRES

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ABSTRACT: Cattle breeding programmes need objective criteria in order to evaluate and subsequently improve production systems. This work uses a logistic growth curve model for evaluating sires based on their progeny weight measured repeatedly over time. The parameters of the curve are described as a linear function of fixed and random effects. A Bayesian approach is used for the estimation. Analysis of the weights recorded on animals of the Nelore breed shows that growth curve models with fixed and random effects can be useful to evaluate and selecting sires.

Key words: beef cattle, logistic model, longitudinal data, random effects

CURVAS DE CRESCIMENTO DE EFEITOS MISTOS NA AVALIAÇÃO DE TOUROS NELORE

RESUMO: Programas de melhoramento de bovinos necessitam critérios objetivos para avaliar e subsequentemente melhorar o sistema de produção. Este trabalho faz uso de um modelo de curva de crescimento logístico para avaliar touros com base nos pesos de suas progênes registradas repetidamente ao longo do tempo. Os parâmetros da curva são descritos como funções lineares de efeitos fixos e aleatórios. Uma abordagem Bayesiana é considerada para a estimação dos parâmetros. Análise de dados de pesos de animais da raça Nelore mostra que os modelos de curva de crescimento com efeitos fixos e aleatórios podem ser úteis para avaliar e selecionar touros.

Palavras-chave: gado de corte, modelo logístico, dados longitudinais, efeitos aleatórios

INTRODUCTION

When considering the importance of selection in genetic breeding programs, several criteria have been proposed as measures of success. Weight measured over time (or at given ages) or gains of weight in a particular period are traits generally used for modelling growth performance of beef cattle.

To analyse growth data of beef cattle, non-linear functions such as Richard's, von Bertalanffy, Brody's, Gompertz's and Logistic growth curves, which describe the weight-age relationship, have been used with success (Fitzhugh Jr, 1976; DeNise & Brinks, 1985; Perotto et al., 1992; Blasco et al., 2003; Freitas, 2005; Forni, 2007). Another approach is to adjust weights to standard ages and use uni- or multivariate analyses to estimate genetic parameters and to predict breeding values for weight at these ages (Kirkpatrick & Heckman, 1989; Eler et al., 1995). A recent approach is to consider the random regression (Jamrozik & Schaeffer, 1997; Olori et al., 1999; Meyer, 2000; Albuquerque & Meyer, 2001) to model traits that are recorded repeatedly over time such as

weight at different ages. Robert-Granié et al. (2002) extended the random regression model to a more general class of models termed heteroskedastic random regressions. This class of models assumes that all variances of random effects can be heterogeneous. In addition, linear mixed models have also been considered by Albuquerque & Fries (1998) for modelling average daily gain of weight and the number of days to gain a specified weight from birth to weaning, in order to propose two animal selection criteria.

The aim of this work is to evaluate the breeding values of sires using a statistical model that is based on weights measured repeatedly on their progeny. This model is a version of the growth curves (Fitzhugh Jr, 1976) within a non-linear mixed-effects model framework (Pinheiro & Bates, 2000), with allowance for correlations between repeated records and also among weights measured on progeny generated by the same sire. Associations are allowed in the model through shared random effects. Thus, instead of considering only the overall growth curve, we assume that there is a separate growth curve for each progeny. A Bayesian approach is used for the estimation and the pa-

parameters drawn from the posterior densities are used to summarize the breeding values of the sires.

MATERIAL AND METHODS

Data used of this paper belong to 6,591 progeny of 63 Nellore sires and 5,251 dams, born during springs between 1993 and 1998 in a single herd. Nellore is a Zebu breed from India that has become predominant in Brazil due to easy care and adaptability. Beef production is the main economic destination of this cattle breed in Brazil. The number of progeny per sire varied from 3 to 460 and per dam from 1 to 5. Factors recorded and considered in the statistical analysis were sex, progeny birth year, management group, and age of the dam at birth of calf, which varied from 2 to 17 years. Approximately 65% of the progeny were female. About 10% of them were born in 1993, 3% in 1994, 13% in 1995, 28% in 1996, 34% in 1997, and 12% in 1998. A total of 102 management groups were observed, i.e., groups of progeny sharing the same permanent environment (quality of pasture, climate, amongst others) from birth to weaning. The minimum and maximum numbers of progeny per group were 11 and 176, respectively. All progeny were weighted individually at six ages. The first weight was taken at birth and the last after two years (between 740 and 886 days after birth). Descriptive statistics of these weights are shown in Table 2.

The statistical model we assumed to evaluate the sires, the logistic mixed-effects growth curve (MGC) model, is based on non-linear functions and uses the progeny weights recorded at all six ages mentioned as outcome vectors. According to Pinheiro & Bates (2000), interpretability, parsimony and validity beyond the observed range of the data are motivations for using non-linear mixed models rather than a competitor linear model, such as a polynomial, when analysing grouped data. A non-linear model generally uses fewer parameters than a linear model and gives a more parsimonious description of the data. In the logistic MGC model,

which describes the trajectory of the growth of each progeny, the weights are expressed as:

$$W_i(t) = \frac{A_i}{(1 + b_i \exp\{-k_i t\})} + \varepsilon_i \tag{1}$$

where $W_i(t)$ is the observed weight of the progeny i on time t ($i = 1, \dots, n$), A_i is the asymptotic weight for progeny i , commonly interpreted as the adult weight. The parameter b_i is a constant of integration that adjusts for situations where $W(0)$ and/or t_0 are not zero and k_i is a function of the ratio between maximum growth rate and mature size, commonly referred to as maturing index (Fitzhugh Jr, 1976).

Generically, a MGC model can be represented as follows:

$$W_i(t) = f(A_p, b_p, k_p, t) + \varepsilon_i$$

where f is a non-linear function expressed in terms of A_p, b_p, k_p and t with A_p, b_p and k_p expressed as linear combinations of fixed and random effects. Extensions to other growth curves are straightforward as can be seen from Table 1.

Thus, calling \mathbf{A}, \mathbf{b} , and \mathbf{k} the vectors of the growth curve parameters A_p, b_p and k_p of all progeny, we considered $\mathbf{A} = \mathbf{X}\boldsymbol{\beta}_a + \mathbf{S}\mathbf{z}_a + \boldsymbol{\varepsilon}_a$, $\mathbf{b} = \mathbf{X}\boldsymbol{\beta}_b + \mathbf{S}\mathbf{z}_b + \boldsymbol{\varepsilon}_b$ and $\mathbf{k} = \mathbf{X}\boldsymbol{\beta}_k + \mathbf{S}\mathbf{z}_k + \boldsymbol{\varepsilon}_k$ where $\boldsymbol{\beta}_a, \boldsymbol{\beta}_b$ and $\boldsymbol{\beta}_k$ are the fixed effects, and $\mathbf{z}_a, \mathbf{z}_b$ and \mathbf{z}_k are the sire random effects for the parameters of the growth curve, \mathbf{X} and \mathbf{S} are incidence matrices, and $\boldsymbol{\varepsilon}_a, \boldsymbol{\varepsilon}_b$ and $\boldsymbol{\varepsilon}_k$ are vectors of random errors. As in a Bayesian approach prior distributions are assigned to all unknown quantities, we consider all fixed and random effects to have normal

Table 1 - Expression of the Gompertz, Richard and Brody growth curves where A_p, b_p, k_p are the same as described previously and m is a shape parameter.

Growth curves	MGCM
Gompertz	$W_i(t) = A_i \exp\{-b_i \exp(-k_i t)\} + \varepsilon_i$
Richard	$W_i(t) = A_i [(1 - b_i \exp(-k_i t))]^m + \varepsilon_i$
Brody	$W_i(t) = A_i [1 - b_i \exp(-k_i t)] + \varepsilon_i$

Table 2 - Means, standard deviations (sd), minimum, maximum and medians of the weights recorded at six ages on 6591 Nellore progeny.

Progeny age	Mean	sd	Minimum	Maximum	Median
Days	----- kg -----				
at birth	30.33	3.47	15	46	30
67 to 138	88.46	17.07	43	171	86
134 to 236	176.50	23.70	100	289	175
282 to 454	304.50	38.18	202	459	300
380 to 672	372.60	43.32	270	570	366
740 to 886	435.70	47.18	346	646	427

prior distributions. For the random errors ε_i in Model (1), for instance, we assumed $N(0, \sigma_\varepsilon^2)$. Also we assumed ε_{ia} , ε_{ib} and ε_{ik} normal distributed with mean 0 and variance σ_a^2 , σ_b^2 and σ_k^2 , respectively. For the hyperparameters τ_a , τ_b , τ_k and τ_e where $\sigma_a^2 = 1/\tau_a$, $\sigma_b^2 = 1/\tau_b$, $\sigma_k^2 = 1/\tau_k$, and $\sigma_e^2 = 1/\tau_e$ we assumed $G(10^{-4}, 10^{-4})$ where $G(\alpha, \lambda)$ indicates a gamma distribution with mean α/λ and variance α/λ^2 . For all β parameters we assumed $N(0, 10^3)$. The sire random effects were also assumed normally distributed. A program for the BUGS software (Spiegelhalter et al., 2000) was constructed to perform the analyses using the Markov Chain Monte Carlo (MCMC) method (Gilks et al., 1996). Posterior distributions of the ranks are used as a guide to help sire evaluation. To obtain the posterior mean of the parameters, a chain of 50,000 iterations was carried out after discarding the first 5,000.

From Model (1) breeders can also estimate the time to reach a desired weight for each progeny i ($i = 1, \dots, n$), as for instance 190 kg, by:

$$T_i = \frac{\log(190b_i) - \log(A_i - 190)}{k_i}, \quad A_i > 190 \quad (2)$$

Sires with few progeny can be used in the proposed model but conclusions about them should be drawn with care.

RESULTS AND DISCUSSION

Descriptive statistics for the weights measured at six ages from birth to past-two years are presented in Table 2. Values for weight at birth varied from 15 to 46 kg with a mean of 30.33 (sd = 3.47). For weights measured between ages 740 and 886 days (past-two years) the mean was 435.7 kg (sd = 47.18). As expected the variance of the weights increases over time.

In order to evaluate the sires we need to focus on particular parameters of the growth curves. Both A , the asymptotic adult weight, and k the maturing index, could be considered since each has a clear biological meaning. Although negative genetic correlation is usually expected between parameters A and k (Notter et al., 1990; Barbato, 1991; Santoro et al., 2005), positive genetic correlation has been found for Nelore, polled Nelore and Guzerá cattle when logistic, von Bertalanffy and Richards growth curves are used (Santoro et al., 2005; Forni, 2007; Forni et al., 2007). In this work we focussed attention on the parameter A . Sire selection based on this parameter has in mind an increasing of the overall weight (or of the growth rate) up to the age in which information are available (around 2.5 years in this work). Table 3 shows a summary of the posterior densities for some of the significant fixed effects considered in the linear function of the parameter A of the growth curve. The male effect was 47.22 with 95% credible interval of (45.13; 49.13) (Table 3). Males gained weight faster than females since the estimate is positive and the interval does not include the zero value. Similarly, a second order polynomial effect was observed for the age of dam at birth of calf (the age and age² coefficients). This means that the progeny weight gain increases until a certain age of the cow, and then decreases. The effects of progeny birth year, and management group (not shown in table) were also significant showing that both affected the progeny weight gain. The posterior mean of the sire variance was $\sigma_s^2 = 69.72$, which gives evidence of a substantial sire effect. A summary of the estimates of the logistic growth curve parameters A_i , b_i , and k_i ($i = 1, \dots, 6591$) is shown in Table 4. Given that for Zebu breeds the adult weight is reached between four and six years, estimates of A_i might be slightly underestimated in this study since weights are available only up to 2.5 years.

Table 3 - Means, standard deviations, medians and credible intervals of the posterior densities of some fixed effects associated with the parameter A of the growth curve where age means dam age at calf birth and byear the birth year of the progeny.

Parameter	Posterior Mean	Standard Deviation	Posterior Median	Credible Interval	
				2.5%	97.5%
$\beta_{a: \text{sex (male)}}$	47.22	0.976	47.23	45.13	49.13
$\beta_{a: \text{age}}$	4.286	0.807	4.374	2.686	5.693
$\beta_{a: \text{age}^2}$	-0.272	0.048	-0.276	-0.355	-0.176
$\beta_{a: \text{byear (1994)}}$	-41.72	20.59	-41.62	-82.06	-2.504
$\beta_{a: \text{byear (1995)}}$	-36.91	6.90	-37.19	-50.22	-22.57
$\beta_{a: \text{byear (1996)}}$	-32.26	6.26	-32.61	-44.24	-16.82
$\beta_{a: \text{byear (1997)}}$	-32.93	6.17	-33.26	-44.52	-19.96
$\beta_{a: \text{byear (1998)}}$	-42.99	6.35	-43.04	-55.06	-26.98

Estimates of the sire effect and credible intervals are presented in Figure 1. As the sire effect was additively incorporated, high and positive values of z_{aj} ($j = 1, \dots, 63$) are desirable. Sires with few progeny (e.g. sire 28) have wider intervals. Clearly superior sires were sires 1, 5, 20, 26, 29, 37, 58, and 60 (Figure 1). From the posterior distribution of the ranks, which provide the uncertainty associated with the rank attributed to each sire, these sires presented a high probability to be classified as good performers (Figure 2). The opposite can be seen, for instance, for sires 28, 30, 31, 54 and 57 where a high probability was observed for low ranks. For these sires a poor performance would be expected from their progeny, in the sense that the progeny demands more time to gain weight. Some sires like 12, 38 and 52, presented high probabilities to be classified in intermediate ranks, so that median performance would be expected from their progeny.

For illustration purposes, two sires among the 63 in this study were chosen such that the first has a

large posterior value of z_a and the second a small posterior value of z_a . Sires 1 ($z_{a1} = 11.95$) and 57 ($z_{a57} = -12.06$) were chosen. Substituting in (2) the estimated A_i values ($i = 1, \dots, 6591$), obtained in the process of estimation, times, T_p , to reach 190 kg were estimated. Progeny from sire 1 would reach 190 kg faster than those from sire 57 (Table 5). Male progeny from sire 1 took a mean time of 194.9 days (sd = 17.8) to reach 190 kg whereas male progeny from sire 57 took 205.6 days (sd = 17.3). Female progeny took longer times to reach 190 kg. Females generated by sire 1 presented a mean time of 236.6 days (sd = 19.3) whereas females from sire 57, 244.7 days (sd = 19.1). Not more than 161 days for males and 195 days for females were needed for the progeny generated by sire 1 to reach 190 kg. All male and female progeny from sire 1 took less than 237 and 285 days, respectively, to reach 190 kg. Hence, progeny from sire 1 presented better performance than those from sire 57.

For the male and female progeny from the two sires 1 and 57 we estimated the logistic growth curves

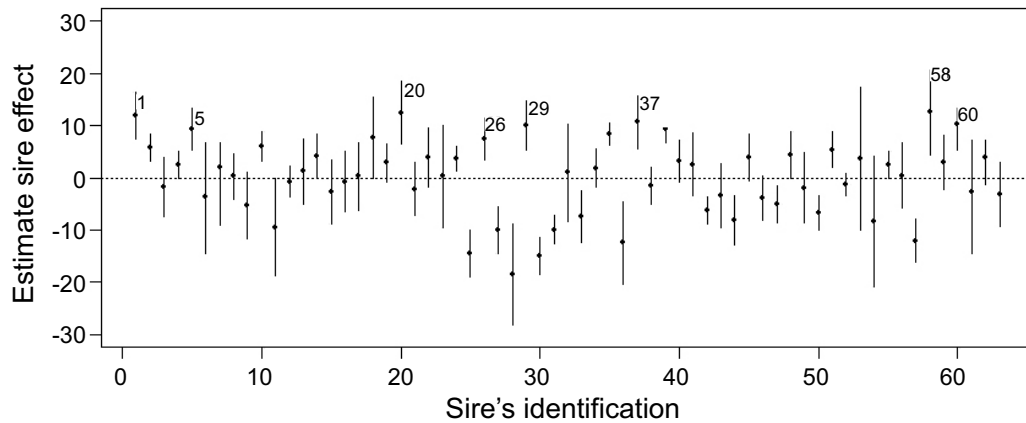


Figure 1 - Posterior means and 95% credible intervals of the z_{aj} estimated sire effects associated with the parameter A of the growth curve.

Table 4 - Summary of the estimates of the parameters A_i , b_i , and k_i of the fitted logistic growth curves.

Parameter	Minimum	Median	Mean	Maximum	Standard deviation
A_i	342.4	418.1	425.3	635.8	41.80
b_i	6.64	8.38	8.42	13.32	0.64
k_i	0.0066	0.0086	0.00886	0.01483	0.0005

Table 5 - Means, standard deviations (sd), minimum and maximum of the time in days estimated to reach 190 kg for the male and female progenies from sires 1 and 57.

Sire	Progeny by sex	Mean time	sd	days	
				Minimum	Maximum
1	Male (n = 24)	194.9	17.8	161.3	236.9
	Female (n = 32)	236.5	19.3	194.7	284.8
57	Male (n = 20)	205.6	17.3	178.6	251.3
	Female (n = 59)	244.7	16.8	202.9	311.8

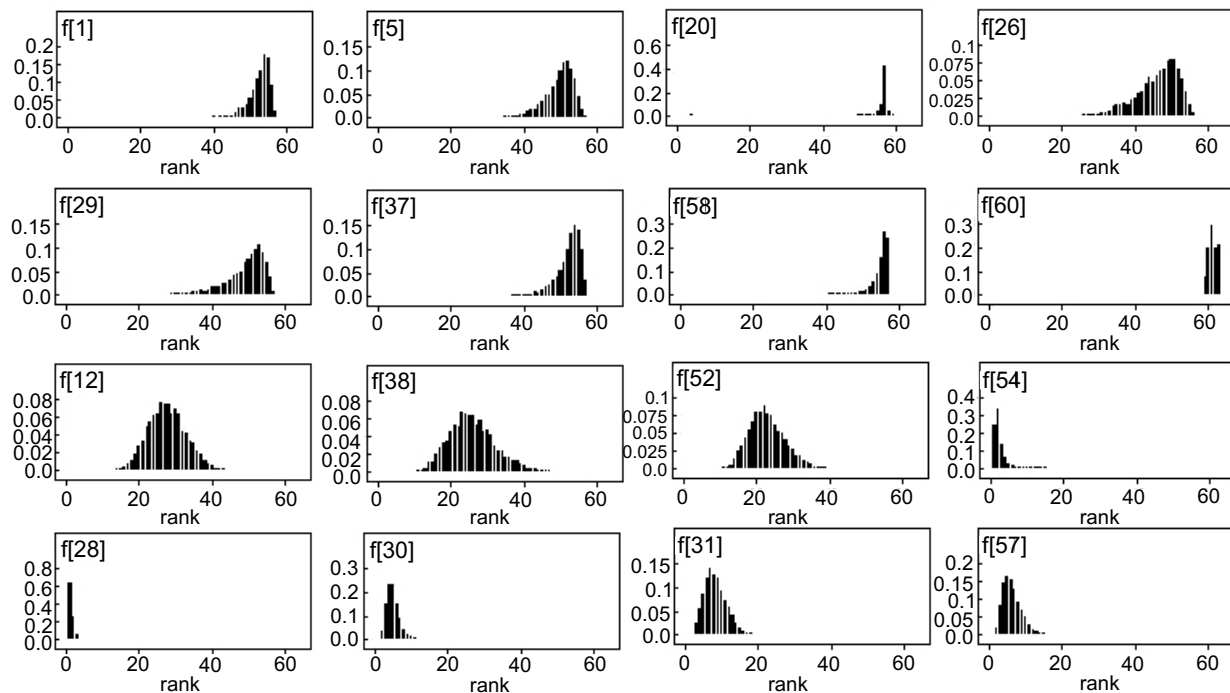


Figure 2 - Posterior distribution of the rank for some of the 63 Nellore sires under study.

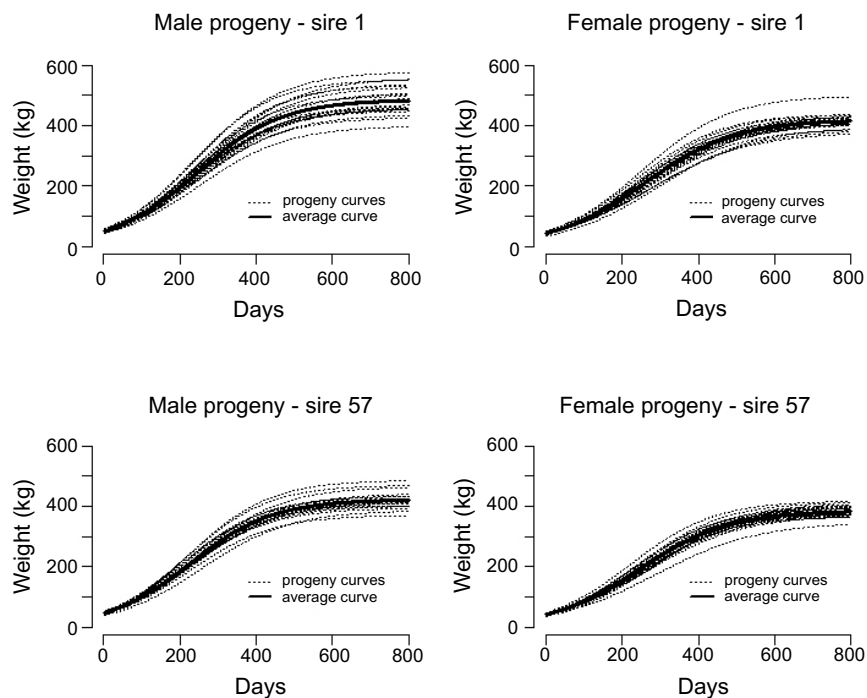


Figure 3 - Mixed logistic growth curves (---) fitted for all progeny of sire 1 (24 males and 32 females) and all progeny of sire 57 (20 males and 59 females) and associated average growth curves (—).

using the Model (1). The progeny from sire 1 presented growth rates higher than those from sire 57 (Figure 3). Comparison among progeny from a specific sire (Figure 3) can be useful not only to evaluate sires which generate progeny of high growth rate, but also to evalu-

ate if the growth rates are homogeneous, i.e. if progeny from a specific sire gain weight uniformly over time. Estimates of the number of days to reach the specific weight of 190 kg or any other weight of interest within the observed range can be obtained (Figure 3).

CONCLUSIONS

Growth curve models with fixed and random effects constitute a valuable data analysis tool for beef cattle breeding since they can be used as an additional and useful tool for selecting sires in cattle breeding programs.

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