

PLOT SIZE AND EXPERIMENTAL UNIT RELATIONSHIP IN EXPLORATORY EXPERIMENTS

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ABSTRACT: Quality of field experiments data dependent upon adequate experimental design. This study investigated the relationship between the size of the basic unit for exploratory experiments and the optimum plot size, and the experiment precision with potato crop. Tuber yield of 24 rows, 144 mounds, were used to arrange experimental units of one, two, three, four, six, eight and 12 mounds. Tuber yield data of different mounds were used to arrange plot sizes with different numbers of basic units. The model $CV(x)=A/X^B$ was adjusted, in which $CV(x)$ was the coefficient of variation among plots with different numbers of basic units. Optimum plot size was estimated with the modified maximum curvature method of the function $CV(x)=A/X^B$, allowing completely random design. Experimental precision was estimated through the Hatheway method for different experimental arrangements. Based upon the modification of the maximum curvature method, for a fixed total experimental area, experimental unit size of potato initial experiments affects optimum plot size estimations with the same experimental precision and variable number of treatments.

Key words: *Solanum tuberosum* L., experiment arrangements, quality control, experimental precision

RELAÇÃO DO TAMANHO ÓTIMO DE PARCELA COM O TAMANHO DA UNIDADE BÁSICA NOS ENSAIOS EM BRANCO

RESUMO: A qualificação dos dados de experimentos de campo é obtida, em grande parte, pela escolha adequada do delineamento experimental. Este estudo foi realizado com o objetivo de investigar a relação do tamanho da unidade básica de ensaios em branco com o tamanho de parcela e com a precisão de experimentos com a cultura de batata. Os dados da produção de tubérculos de batata, obtidos de 24 linhas com 144 covas, foram usados para planejar unidades básicas (UB) de 1, 2, 3, 4, 6, 8 e 12 covas. Com os valores da produção das UB de diferentes tamanhos foram planejadas parcelas de diferentes números de UB. Foi ajustado o modelo $CV(x)=A/X^B$, em que $CV(x)$ é o coeficiente de variação entre as parcelas com X UB de tamanho. A estimativa do tamanho ótimo de parcela foi procedida pelo método da máxima curvatura modificada da função $CV(x)=A/X^B$, considerando o delineamento inteiramente ao acaso, e a precisão foi estimada pelo método de Hatheway para diferentes configurações do delineamento experimental. Concluiu-se que o tamanho da unidade básica, usada nos ensaios em branco com batata, afeta a estimativa do tamanho ótimo de parcela pelo método da máxima curvatura modificada, sem alterar a precisão experimental, independente do número de tratamentos, mantendo fixa a área total do experimento.

Palavras-chave: *Solanum tuberosum* L., planejamento experimental, controle de qualidade, precisão experimental

INTRODUCTION

The search for more suitable, productive genotypes demands additional effort from researchers, mainly in the areas of genetic improvement and experimental techniques. With the use of information about features of the experimental area and the planted crop, it is possible to reduce the experimental error. The magnitude of the experimental error, can be affected by several sources of error, such as: type of treatment; presence of pests, diseases and weeds; interplot and intraplot competition; the

heterogeneity of the experimental material, and the heterogeneity of the experimental units (Storck et al., 2000; Ramalho et al., 2000).

The problem of heterogeneity in experimental units can be solved by choosing a suitable experimental design, using covariance analysis and establishing suitable plot size and shape (Gomez & Gomez, 1984; Steel et al., 1997). In general, mainly in limited areas, it is preferable to use a greater number of replications, allied to smaller plots, provided they have the practical matching with the treatments to be applied (Storck et al., 2000).

The establishment of experimental plot size and shape could be achieved on an empirical basis, from the experience of researchers who have already worked with the crop to be studied, and their knowledge on the experimental plot. Nevertheless, there are several methods that provide such information in a more objective way, by optimizing costs, arranging the number of replications and also making use of data from field experiments (Kock & Rigney, 1951; Pimentel Gomes, 1988) or by using exploratory or uniformity experiments (Smith, 1938; Keller, 1949; Hatheway & Williams, 1958; Hatheway, 1961; Lessman & Atkins, 1963; Meier & Lessman, 1971; Thomas, 1974; Hach & Morales, 1976; Lugo, 1977; Vallejo & Mendoza, 1992; Oliveira & Estefanel, 1995; Simplicio et al., 1996; Muniz et al., 1999; Viana et al., 2002a; 2002b; Silva et al., 2003; Mello et al., 2004). The aim of this study is investigating the relationship between the size of the basic unit for exploratory experiments and the optimum plot size, and the experiment precision with potato crop.

MATERIAL AND METHODS

Potato crop data were collected during the agricultural year of 2003, in Júlio de Castilhos RS, Brasil (29°12'S, 53°41'W; altitude 490 m). The soil in the area is a Dystrophic Dark Red Latosol Soil.

Tubers of 3,456 mounds were harvested on 28/11/2003 and 01/12/2003 from one area of three hectares, cultivated with potato cultivar macaca for the multiplication of seeds, planted on August 10, 2003, and weighed with a portable, digital balance. The exploratory experiments were set on an area of 0.0691 ha, in which the mounds were identified according to the number of rows (1 to 24) and columns (1 to 144). Rows were spaced 0.8 m, with about four mounds per linear meter or 0.2 m² per mound, a density of 50,000 mounds ha⁻¹. The area was fertilized, provided with recommended cultivar and phytosanitary treatment, and irrigated three times during the cycle. Twelve days before the harvest the desiccant Gramoxone was applied.

Considering the potato tuber yield results from each mound, different plot sizes and shapes were arranged by adding the results of neighboring mound towards the (L) row and the (C) column. Different Basic Unit (BU) sizes were then arranged by adding one, two, three, four, six, eight and 12 mounds in the (P) row. Later, with each BU size (P), different plot types were arranged. The number of replications (N) for the plots arranged with an L (ranging from 1, 2, 3, 4, 6) and C (ranging from 1, 2, 3, 4, 6, 8, 12), with L≤C restriction, was given by the formula $N=3456 / PLC$; $X=LC$ is the number of BU per plot. Statistics were calculated for each P value: $M(x)$ = the average of plots with X BU; $VU(x)$ = the variance per BU, calculated between plots with X BU, given by the

expression $VU(x) = V(x)/X^2$, where $V(x)$ is a variance between the plots of X BU size; $CV(x)$ = the variance coefficient between the plots of X BU size.

The Smith (1938) soil heterogeneity index “b” was estimated as being the linear regression coefficient of the function $VU(x) = V_1/X^b$, transformed to a logarithm with weighted estimate by the degrees of freedom associated with the types of planned plots (Steel et al., 1997). To estimate the optimum plot size, the modified maximum curvature method was used (Meier & Lessman, 1971). This method uses the function $CV(x) = A/X^b$, in which the A parameter estimates the value of the plot variation coefficient of an BU of size and B estimates the curvature between $CV(x)$ and the plot size. Similarly, in the function of Smith (1938), A and B values are estimated by the least squares method using the logarithm transformation, and the weighting as degrees of freedom. The maximum curvature point of the function $CV(x) = A/X^b$ is the optimum plot size (X_0), thus:

$$X_0 = \exp \left\{ \left[\frac{1}{2B+2} \right] \ln \left[\frac{A^2 B^2 (2B+1)}{B+2} \right] \right\}$$

To study the relation between the optimum plot size transformed in area ($m^2=0.2 \cdot P \cdot X_0$) and the respective BU sizes, in number of mounds (P), several pre-established functions were tested in the statistical SAEG package (Ribeiro Jr., 2001). The selection criterion of the function was applied by the significance of parameters estimates (t test) and by the value of determination coefficient.

The true difference between two treatments (measured as percent of the mean) (d) was estimated using the method of Hatheway (1961), formula $d = \sqrt{2(t_1 + t_2)^2 A^2 / JX_0^b}$, where “J” is the number of replications for different plot sizes (X_0) in numbers of BU; “A” and “b”, respectively, are estimates from the coefficient of variation between the plots of an BU and soil heterogeneity index; t_1 is the significant value of t in the test of significance (bilateral to 5%); t_2 is the value of t in table corresponding to 2(1-Pr), where Pr is the probability of obtaining a significant result (0.80). The t distribution values were obtained using DF degrees of freedom taking into account the completely randomized design, i.e. $DF=I(J-1)$, where I is the number of treatments and J the number of replications. The number of treatments was arbitrated in nine, 12, 18 and 24, and the number of replications was calculated by $J=3456/(PX_0I)$.

The calculations were accomplished by using the following statistical packages: scientific software – NTIA, developed by the Centro Tecnológico para Informática (EMBRAPA, 1997), the GENES program (Cruz, 2001), the SAEG program (Ribeiro Jr., 2001) and a program in the FORTRAN language, specifically used to calculate averages and plot variances of different sizes.

RESULTS AND DISCUSSION

The average tuber yield was $0.8464 \text{ kg mound}^{-1}$, with standard deviation (SD) = $0.30158 \text{ kg mound}^{-1}$, and coefficient of variation (CV) = 35.63%. From these values, it was estimated an yield of $42,32 \text{ t ha}^{-1}$, a very high value in comparison to the average Brazilian crop yield of potato for consumption (15 t ha^{-1}), and to the average crop yield in Rio Grande do Sul (8 t ha^{-1}) (Pereira & Daniels, 2003). In Brazil, top potato farmers harvest between 30 and 35 t ha^{-1} , and in European countries, the average crop yield nears 40 t ha^{-1} (Filgueira, 2002). Despite the existing variability between mounds (CV=35.63%), the area presented clear homogeneity of horizontal spacing between plants regarding attack of insects, diseases and weeds, which could adversely affect the use of data for the proposed objective.

There were no variations in the estimates of the function $CV(x) = A/X^B$ parameters, and in the determination coefficient (r^2), possibly related to the number of mounds per BU (Table 1). The increase in the number of mounds per BU up to four caused a reduction in the X_0 value, oscillating in the larger values. Taking into account the amplitude X_0 equal to 1.25 BU, and the average equal to 3.09 BU, no dependency of X_0 seems to be evident with the number of mounds per BU. The question is if the use of a greater number of mounds per BU alters, almost linearly (quadratic and potential), the area of the optimum size of the estimated plot using the maximum curvature method (Figure 1). The relation between the number of mounds per BU and the area of the optimum plot size is almost perfect, as the determination coefficient for both models of Figure 1 is close to one, and it is therefore evident the relationship between the area of the optimum plot size estimated by the method of maxi-

imum curvature modified, and the BU size in the uniformity tests. Therefore, the estimate methods of the optimum plot size that use experimental data with treatment effects, usually with plots greater than a mound can overestimate the optimum plot size.

The average of the optimum plot size (3.09 BU) is contained in the diversity of results with the potato crop obtained by Oliveira & Estefanel (1995). However, to compare the obtained results, the optimum plot size should be transformed in number of mounds. In this case, Oliveira & Estefanel (1995) obtained a four to 90 variation of mounds for different methods and times, since they used a three mound BU.

Some studies conducted with potato crops, such as those of Brown (1987) and Bearzoti & Pinto (1996), discussed the plot size based on the stage of the improvement program carried out with that crop, since initial phases, when there are too many genotypes to be tested and few tubers available, smaller plots should be prioritized, whereas in the following stages, with fewer genotypes to be tested and the need for detecting little meaningful differences between averages of clones, one should use larger plots, which could be associated with the method to be used, and also with the BU size.

Considering that the soil heterogeneity index (b) of the function $VU(x) = V_1/X^b$ is equal to 2B, B of the function $CV(x) = A/X^B$ (Thomas, 1974), it is verified that the b values with average equal to 0.36 ($b=2B=2*0.18$) vary slightly in relation to the number of mounds per BU and is regarded as low, which indicates the need for small plots. The low value of b is likely due to the size reduction of stolons, obtained from the culture development, whose phenotypic selection is made during the harvest of the first generation of clones (Bisognin, 2003), in which the tuber yield is very concentrated around the seed-potato, along with the fertilizer applied in the row.

Table 1 - Estimates of the function $CV(x) = A/X^B$ parameters, determination coefficient (r^2) and the optimum plot size (X_0), obtained by the modified maximum curvature method, to the variable potato tuber yield, using basic units (BU) with different number of mounds.

BU	A	B	r^2	$*X_0$
1 mound	31.0870	0.1966	0.8101	3.7529
2 mounds	27.9914	0.1809	0.7910	3.2356
3 mounds	27.0575	0.1816	0.7991	3.1530
4 mounds	24.3511	0.1517	0.7616	2.5017
6 mounds	25.5379	0.1856	0.8247	3.0520
8 mounds	23.9731	0.1666	0.9238	2.6615
12 mounds	24.5506	0.2162	0.8823	3.2971
Average	26.3641	0.1827	0.8275	3.0934

$$*X_0 = \exp\left\{\left[\frac{1}{2(B+2)}\right] \ln\left\{\frac{A^2 B^2 (2B+1)}{(B+2)}\right\}\right\}$$

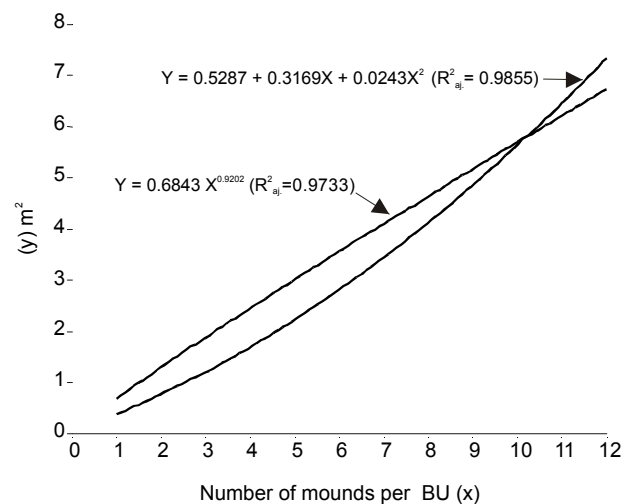


Figure 1 - Relationship between the area, estimated by the modified maximum curvature method, and the number of mounds per basic unit.

Table 2 - True difference between two treatments (measured as percent of the mean) (d) obtained by the Hatheway method, and the number of repetitions (J) for different plot sizes (X_0), in number of basic units (BU), comprised of two and four mounds and different amounts of treatments (I), for the potato tuber yield.

Mounds per BU	X_0	I = 9		I = 12		I = 18		I = 24	
		J	d	J	d	J	d	J	d
2	1	192	8.07	144	9.32	96	11.42	72	13.18
	2	96	10.08	72	11.64	48	14.25	36	16.45
	3	64	11.47	48	13.25	32	16.22	24	18.73
2	4	48	12.58	36	14.52	24	17.79	18	20.54
	6	32	14.32	24	16.54	16	20.25	12	23.39
	8	24	15.70	18	18.13	12	22.21	9	25.64
2	12	16	17.88	12	20.64	8	25.28	6	29.19
	1	96	9.78	72	11.30	48	13.84	36	15.98
	2	48	12.47	36	14.40	24	17.64	18	20.37
4	3	32	14.37	24	16.60	16	20.33	12	23.47
	4	24	15.90	18	18.35	12	22.48	9	25.96
	6	16	18.32	12	21.15	8	25.91	6	29.91
4	8	12	20.26	9	23.39	6	28.65	-	-
	12	8	23.35	6	26.96	4	33.02	3	38.13

Therefore, it is very likely that, in the case of potato crops, the heterogeneity of the production is mainly due to the heterogeneity in the fertilization (of rows) and the seed-potato size, and in less intensity, due to the variation in the original conditions of the soil, such as fertility, moisture, planting depth, among others.

To facilitate planting depth, fertilization process and consequent culture treatments, potatoes are cultivated in mounds arranged in lines, in which the latter is more homogeneous. Therefore, when experimental blocks are inserted, the blocking should be constituted in the line itself, and the plots formed would be narrow (one line), forming narrow blocks (one line) with the length based on the number of treatments.

When applying the Hatheway (1961) method using BU of two and four mounds with variation in the number of replications (J), in the optimum plot size (X_0), and in the number of treatments (I), in a way to plan experiments so as to occupy the entire area evaluated, for the same number of mounds per parcel (number of BU multiplied by the number of mounds per BU), and the same number of replications, the precision (d) is similar, regardless of the number of treatments (Table 2). Therefore, the determination of the optimum plot size in experiments should be seen in the set, by combining J and I, and the area available for a pre-established precision. Smaller plots with a greater number of repetitions in the same total area result in greater precision, but with a greater number of studies (evaluations), whose capacity for executing operations should be well evaluated before the experiment installation.

In general, with basic units of four mounds, in order to obtain d around 20%, it should work with, for example, I=18, J=16 and $X_0=3$ (Table 2). Yet, the same precision can be obtained with basic units comprised of two mounds, by choosing I=18, J=16 and $X_0=6$. Therefore, when reducing the number of mounds per basic unit by half, the plot size estimate is doubled in number of basic units in order to achieve compensation and maintain the accuracy. Therefore, it is concluded that the BU size, used in the exploratory experiments with potato, affects the estimate of the optimum plot size by the maximum modified curvature method, without altering the experimental accuracy, regardless of the number of treatments and maintaining total area of the experiment fixed.

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